

# Accelerated Testing with Recurrent Events

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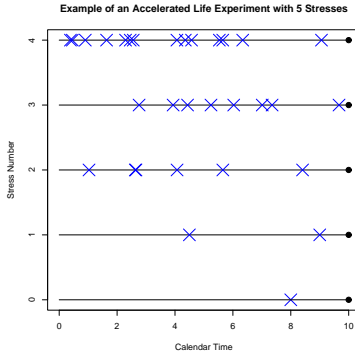
# Outline

- Introduction
- Model
- Methods
- Optimal Design
- Future Research

# Accelerated Life Testing

- Accelerated Life Testing (ALT) induces failures through high stress (temperature, humidity, usage etc.)
- Stresses are joined through a link function (inverse power, log-linear, Arrhenius, etc.)
- Results used designing warranties, system comparison, and failure detection

# Example Accelerated Life Testing



**Figure:** Graphical view of an ALF with 5 Stresses with one site for each stress.

## Previous Literature

- Chernoff ('62) - simultaneous testing with type I censoring and successive testing with complete data under HPP
- Nelson and Kielpinski ('76); Nelson and Meeker ('78)- type I censored simultaneous testing for Normal, Log-Normal, Weibull, and extreme value distributions

# Problem Statement

## Goals:

- 1 Parameter Estimation
- 2 Hypothesis tests and CI estimation on the percentile of the assumed life distribution at the operating stress level
- 3 Develop optimum ALT test plans

# Notation

- 1  $i = \text{stress}, j = \text{site}, l = \text{event}.$
- 2 Counting Process

$$N_{ij}^{\dagger}(s) = \sum_{l=1}^{\infty} I(\mathbf{S}_{ijl} \leq \mathbf{s}, \mathbf{S}_{ijl} \leq \tau_i)$$

- 3 At-risk indicator

$$Y_{ij}^{\dagger}(s) = I\{\tau_i \geq s\}$$

- 4 Backward recurrence time

$$\varepsilon_{ij}(w) = w - S_{iN_{ij}^{\dagger}(w-)}$$

# Basic Assumptions

- Generally the distribution we are considering is:

$$T_{ijl}|x_i \xrightarrow{iid} \Lambda_0 \left[ \frac{t}{\theta(x_i; \beta)}; \alpha \right] \quad \text{where} \quad \log[\theta(x_i; \beta)] = \beta_0 + \beta_1 x_i$$

- Assume  $X_0, X_1, X_2, \dots, X_p$ , possible stresses:
  - $X_p \equiv X_H$  denotes the highest stress
  - $X_0$  is the design stress level
- For the  $i^{\text{th}}$  stress we have  $m_i$  systems with:

$$m_0, m_1, m_2, \dots, m_p \quad M = \sum_{i=0}^p m_i$$

# AFT Model

- The compensator for the  $i^{\text{th}}$  stress and  $j^{\text{th}}$  site will be:

$$A_{ij}^{\dagger}(s; \alpha, \beta_0, \beta_1) = \int_0^s Y_{ij}^{\dagger}(v) \frac{1}{\theta(\mathbf{x}_i; \beta_0, \beta_1)} \lambda_0 \left[ \frac{\varepsilon_{ij}(v)}{\theta(\mathbf{x}_i; \beta_0, \beta_1)}; \alpha \right] dv$$

which is predictable with respect to the common filtration:

$$\mathcal{F}_s = \mathcal{F}_0 \vee \left\{ \bigvee_{i=1}^n \sigma \left\{ \left( N_{ij}^{\dagger}(w), Y_{ij}^{\dagger}(w+), \varepsilon_{ij}(w+) \right) : 0 < w \leq s \right\} \right\}$$

- $M_{ij}^{\dagger}(s) = N_{ij}^{\dagger}(s) - A_{ij}^{\dagger}(s)$  is a zero-mean square integrable martingale for the  $i^{\text{th}}$  stress and  $j^{\text{th}}$  site.

# Maximum Likelihood Formulation

- Following Jacod ('75):

$$L_{ij}(\mathbf{s}; \mathbf{x}_i, \beta, \alpha) = \prod_{i=1}^n \left\{ \left[ d A_{ij}^{\dagger}(\mathbf{s}; \mathbf{x}_i, \beta, \alpha) \right]^{\Delta N_{ij}^{\dagger}(\mathbf{s})} \exp \left[ - A_{ij}^{\dagger}(\mathbf{s}; \mathbf{x}_i, \beta, \alpha) \right] \right\}$$

- Score Process:

$$U(\mathbf{s}; \mathbf{x}_i, \beta, \alpha) = \begin{bmatrix} U_{\beta}(\mathbf{s}; \mathbf{x}_i, \beta, \alpha) \\ U_{\alpha}(\mathbf{s}; \mathbf{x}_i, \alpha, \beta) \end{bmatrix} = \sum_{i=1}^p \sum_{j=1}^{m_i} \int_0^{\mathbf{s}} \mathbf{H} \left[ \frac{\varepsilon_{ij}(w)}{\theta(\mathbf{x}_i; \beta)}; \mathbf{x}_i, \beta, \alpha \right] dM_{ij}^{\dagger}(w; \mathbf{x}_i, \beta, \alpha)$$

## MLE Result:

- Subject to Borgan('84):

$$\sqrt{n}[\hat{\theta} - \theta] \xrightarrow{iid} AN(0, \Sigma^{-1})$$

where  $\hat{\theta} = (\hat{\alpha} \quad \hat{\beta}_0 \quad \hat{\beta}_1)'$  is the Maximum Likelihood estimator of  $\theta = (\alpha \quad \beta_0 \quad \beta_1)'$ .

- A consistent estimate of  $\Sigma$  is:

$$\hat{\Sigma} = \frac{1}{M} \langle \mathbf{U}(\mathbf{s}; \beta, \alpha) \rangle = \frac{1}{M} \sum_{i=1}^p \sum_{j=1}^{m_i} \int_0^s \mathbf{H}_{ij} \left[ \frac{\varepsilon_{ij}(w)}{\theta(x_j; \beta)}; \mathbf{x}_i, \beta, \alpha \right]^{\otimes 2} d\mathbf{A}_{ij}^{\dagger}(w; \mathbf{x}_i, \beta, \alpha)$$

# MLE result: H matrix

- Using the notation that;

$$\gamma(\mathbf{x}_i, \beta) = \nabla_{\beta} \text{Log}[\theta(\mathbf{x}_i; \beta)]$$

$$\pi(\mathbf{t}, \alpha) = \nabla_{\alpha} \text{Log}[\lambda_0(\mathbf{t}; \alpha)]$$

$$\eta(\mathbf{t}, \alpha) = \frac{\partial}{\partial \mathbf{t}} \text{Log}[\lambda_0(\mathbf{t}; \alpha)]$$

We have:

$$\mathbf{H}_{ij} \left[ \frac{\varepsilon_{ij}(\mathbf{w})}{\theta(\mathbf{x}_i; \beta)}; \mathbf{x}_i, \beta, \alpha \right] = \begin{pmatrix} -\gamma(\mathbf{x}_i; \beta) \left[ 1 + \frac{\varepsilon_{ij}(\mathbf{w})}{\theta(\mathbf{x}_i; \beta)} \eta \left( \frac{\varepsilon_{ij}(\mathbf{w})}{\theta(\mathbf{x}_i; \beta)}, \alpha \right) \right] \\ \pi \left[ \frac{\varepsilon_{ij}(\mathbf{w})}{\theta(\mathbf{x}_i; \beta)}, \alpha \right] \end{pmatrix}$$

# Maximum Likelihood Formulation

- After calculations:

$$\langle \mathbf{U}_{ij}(\mathbf{s}; \beta, \alpha) \rangle = \int_0^{\tau} Y_{ij}(w) \mathbf{H}_{ij}(w; \mathbf{x}_i, \beta, \alpha)^{\otimes 2} \lambda_0(w; \alpha) dw$$

# Important Regularity Condition

- Covariation Process:

$$\frac{1}{M} \langle \mathbf{U}(\mathbf{s}; \beta, \alpha) \rangle = \frac{1}{M} \sum_{i=1}^p \sum_{j=1}^{m_i} \int_0^{\tau} Y_{ij}(w) \mathbf{H}_{ij}(w; \mathbf{x}_i, \beta, \alpha)^{\otimes 2} \lambda_0(w; \alpha) dw$$

$$\begin{aligned} \text{as } M \rightarrow \infty &\xrightarrow{d} \sum_{i=1}^p S_i \int_0^{\tau} y(w; \beta, \mathbf{x}_i) \mathbf{G}(w; \mathbf{x}_i, \beta, \alpha) \lambda_0(w; \alpha) dw \\ &= \sum_{i=1}^p \mathbf{Q}(\mathbf{x}_i; \beta, \alpha) \end{aligned}$$

Where  $S_i \equiv \frac{m_i}{M}$ :

$$E[Y_{ij}(s, t)] = y(w; \beta, \mathbf{x}_i) = \bar{F}[t\theta(\mathbf{x}_i)] + \bar{F}[t\theta(\mathbf{x}_i)]\rho[\theta(\mathbf{x}_i)(\tau - t)]$$

[Pena et al. 2001, JASA]

# Optimization Criteria

- Most often the goal in ALT is to estimate a percentile ( $q$ ) of the CDF at the operating stress level:

$$\xi_q(\mathbf{X}_0) = F_0^{-1}(q; \mathbf{X}_0; \beta_0, \beta_1, \alpha) \quad (1)$$

- Minimization Criteria:

$$\text{Min}_{\{S_j, j=1,2,\dots,p\}} \left\{ \text{Var}[\xi_q(\mathbf{X}_0)] \right\}$$

An application of the Delta Method yields:

$$\text{Var}(\xi_q(\mathbf{X}_0)) = B\Sigma^{-1}(s)B^t \quad \text{with} \quad B = \begin{pmatrix} \nabla_{\beta} \xi_q(\mathbf{X}_0) \\ \nabla_{\alpha} \xi_q(\mathbf{X}_0) \end{pmatrix}$$

## Previous Results

- Elfving's ('52) result was then used by Chernoff ('62) to develop the first ALT 'locally optimal' test plans
- Chernoff found that in a two-parameter life-stress relationship only two stresses need to be used, with  $X_H$  being one

## Compromised Plans

- Statistically optimal test plans will always use a minimum amount of stresses, but can are not robust to model misspecifications
- In a two-parameter model Compromised Plans use three or four test stresses

# Future Research

- Develop statistically optimum and compromised plans for exponential, and weibull distributions under a recurrent event framework
- Develop hypothesis tests and confidence intervals for  $\xi_p(\cdot)$
- Develop a nonparametric model under a recurrent event framework

# Thank You

Thanks For Listening