

### STAT 730: Homework 3

1. **An important application of two theorems from Chapter 3 to univariate regression.** Consider the usual regression model from STAT 704:  $\mathbf{y} \sim N_n(\mathbf{X}\boldsymbol{\beta}, \sigma^2\mathcal{I}_n)$  where  $\mathbf{X}(n \times p)$  is full rank,  $n > p$ , and  $\boldsymbol{\beta}(p \times 1)$  are regression effects. The first column in  $\mathbf{X}$  is  $\mathbf{x}_{(1)} = \mathbf{1}_n$ . Let  $\mathbf{X} = [\mathbf{X}_1(n \times (p-k)) \mathbf{X}_2(n \times k)]$  and  $\boldsymbol{\beta}' = (\boldsymbol{\beta}_1((p-k) \times 1)', \boldsymbol{\beta}_2(k \times 1)')$ . We have two models, reduced (R) and full (F):

$$R : \mathbf{y} \sim N_n(\mathbf{X}_1\boldsymbol{\beta}_1, \sigma^2\mathcal{I}_n), \quad F : \mathbf{y} \sim N_n(\mathbf{X}\boldsymbol{\beta}, \sigma^2\mathcal{I}_n).$$

Finally let  $\mathbf{P}_\mathbf{X} = \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'$ ,  $\mathbf{P}_{\mathbf{X}_1} = \mathbf{X}_1(\mathbf{X}_1'\mathbf{X}_1)^{-1}\mathbf{X}_1'$ , and  $\mathbf{P}_{\mathbf{1}_n} = \frac{1}{n}\mathbf{1}_n\mathbf{1}_n'$ . We will construct the F-statistic for testing  $H_0 : \boldsymbol{\beta}_2 = \mathbf{0}$  in the full model.

- (a) Show  $SSR_d = \mathbf{y}'[\mathbf{P}_\mathbf{X} - \mathbf{P}_{\mathbf{X}_1}]\mathbf{y}$  has a scaled  $\chi^2$  distribution under  $H_0$  using Cochran's theorem. Hint: first show  $SSR_d = (\mathbf{y} - \mathbf{X}_1\boldsymbol{\beta}_1)'[\mathbf{P}_\mathbf{X} - \mathbf{P}_{\mathbf{X}_1}](\mathbf{y} - \mathbf{X}_1\boldsymbol{\beta}_1)$  and argue that  $(\mathbf{y} - \mathbf{X}_1\boldsymbol{\beta}_1)$  is a d.m. under  $H_0$ .
  - (b) Show  $SSR_d = SSR_f - SSR_r$  where  $SSR_r = \mathbf{y}'[\mathbf{P}_{\mathbf{X}_1} - \mathbf{P}_{\mathbf{1}_n}]\mathbf{y}$  and  $SSR_f = \mathbf{y}'[\mathbf{P}_\mathbf{X} - \mathbf{P}_{\mathbf{1}_n}]\mathbf{y}$ . This is the difference in SSR from the full and reduced models (it is also the difference in the SSE). To see this, e.g.,  $SSR_f = \sum_{i=1}^n (\hat{y}_i - \bar{y})^2 = \|\hat{\mathbf{y}} - \mathbf{1}_n\bar{y}\|^2 = \|\mathbf{X}\hat{\boldsymbol{\beta}} - \mathbf{1}_n\bar{y}\|^2 = \|(\mathbf{P}_\mathbf{X} - \mathbf{P}_{\mathbf{1}_n})\mathbf{y}\|^2 = \mathbf{y}'[\mathbf{P}_\mathbf{X} - \mathbf{P}_{\mathbf{1}_n}]\mathbf{y}$ , using  $\mathbf{P}_\mathbf{X} - \mathbf{P}_{\mathbf{1}_n}$  idempotent & symmetric.
  - (c) Similarly, show  $SSE_f = \mathbf{y}'[\mathcal{I}_n - \mathbf{P}_\mathbf{X}]\mathbf{y}$  has a scaled  $\chi^2$  distribution under  $H_0$  using Cochran's theorem.
  - (d) Show  $SSR_d$  indep. of  $SSE_f$  using the "partitioning sums of squares", i.e. Craig's theorem.
  - (e) Finally, show  $F^* = \frac{SSR_d/k}{SSE_f/(n-p)} \sim F_{k, n-p}$  under  $H_0$ .
2. Using the previous "big model / little model" result, show that the F-test in the usual ANOVA table for univariate multiple regression has a  $F_{p-1, n-p}$  distribution.
3. Show  $E(\mathbf{x}'\mathbf{A}\mathbf{x}) = \text{tr}\mathbf{A}\boldsymbol{\Sigma} + \boldsymbol{\mu}'\mathbf{A}\boldsymbol{\mu}$  where  $\mathbf{x} \sim (\boldsymbol{\mu}, \boldsymbol{\Sigma})$  is random and  $\mathbf{A}$  is conformable. Hint:  $E\{\text{tr}(\cdot)\} = \text{tr}\{E(\cdot)\}$  and  $\text{tr}\mathbf{A}\mathbf{B} = \text{tr}\mathbf{B}\mathbf{A}$ . Now find  $E(SSR_f)$  when the full model holds; under  $H_a$  in the previous problem.
4. Show  $\frac{\partial \mathbf{a}'\mathbf{x}}{\partial \mathbf{x}} = \mathbf{a}$  and  $\frac{\partial \mathbf{x}'\mathbf{A}\mathbf{x}}{\partial \mathbf{x}} = (\mathbf{A} + \mathbf{A}')\mathbf{x}$ .
5. Let  $\mathbf{A} = [a_{ij}]$ , all elements distinct. Show  $\frac{\partial |\mathbf{A}|}{\partial a_{ij}} = A_{ij}$  where  $A_{ij}$  is the  $ij$ th cofactor of  $\mathbf{A}$ . Hint: use (I) on p. 457 plus a careful argument. Note: the result for symmetric  $\mathbf{A}$  is considerably more difficult.
6. Let  $\mathbf{x}_1, \dots, \mathbf{x}_n \stackrel{iid}{\sim} N_p(\boldsymbol{\mu}, \boldsymbol{\Sigma})$  under the priors  $\boldsymbol{\mu} \sim N_p(\mathbf{m}, \mathbf{V})$  indep.  $\boldsymbol{\Sigma}^{-1} \sim W_p(\mathbf{A}, a)$ . Show
- $$\boldsymbol{\mu} | \boldsymbol{\Sigma}, \mathbf{X} \sim N_p([n\boldsymbol{\Sigma}^{-1} + \mathbf{V}^{-1}]^{-1}[n\boldsymbol{\Sigma}^{-1}\bar{\mathbf{x}} + \mathbf{V}^{-1}\mathbf{m}], [n\boldsymbol{\Sigma}^{-1} + \mathbf{V}^{-1}]^{-1}),$$
- and
- $$\boldsymbol{\Sigma}^{-1} | \boldsymbol{\mu}, \mathbf{X} \sim W_p \left( \left[ \mathbf{A}^{-1} + \sum_{i=1}^n (\mathbf{x}_i - \boldsymbol{\mu})(\mathbf{x}_i - \boldsymbol{\mu})' \right]^{-1}, a + n \right).$$
7. MKB 4.2.5.
8. MKB 4.2.12.