

## STAT 730: Homework 2

1. **MLE of multivariate normal.** We will fill in details of a maximization theorem for  $\Sigma$  in Chapter 4 (p. 105). Let  $\mathbf{A} > 0$  and  $\mathbf{B} > 0$  be symmetric in all that follows.

- (a) Show  $\mathbf{A} > 0 \Leftrightarrow$  all e-values of  $\mathbf{A}$  are positive.
- (b) Show  $\mathbf{A}^{-1} > 0$ .
- (c) Show  $\mathbf{A}^{-1/2}\mathbf{B}\mathbf{A}^{-1/2}$  has the same e-values as  $\mathbf{A}^{-1}\mathbf{B}$ . Hint: let  $\mathbf{v}$  be an e-vector of  $\mathbf{A}^{-1}\mathbf{B}$  and show  $\mathbf{A}^{-1/2}\mathbf{B}\mathbf{v}$  is an e-vector of  $\mathbf{A}^{-1/2}\mathbf{B}\mathbf{A}^{-1/2}$ .
- (d) Show the e-values of  $\mathbf{A}^{-1/2}\mathbf{B}\mathbf{A}^{-1/2}$  are all positive. Hint:  $\mathbf{B} > 0$ .

Thus the e-values of  $\mathbf{A}^{-1}\mathbf{B}$  are all positive by (c) and (d); (a) and (b) are just for fun.

2. **Characteristic function of normal:** Let  $x \sim N(0, 1)$ .

- (a) Show  $E(x^p) = 0$  if  $p$  odd and  $E(x^p) = (p-1)(p-3)(p-5)\cdots 1$  if  $p$  is even. Hint: look at  $\frac{1}{2}E(x^p) = I = \int_0^\infty x^p \frac{1}{\sqrt{2\pi}} \exp\{-\frac{1}{2}x^2\}dx$  and make the change of variables  $y = x^2$ . Then use properties of  $\Gamma(\cdot)$ .
- (b) Now use the Taylor's expansion  $e^x = \sum_{n=0}^\infty \frac{x^n}{n!}$  in  $E\{e^{itx}\}$ , take expectation of each term, and simplify. This works because  $\int_a^b ix^p f(x)dx = i \int_a^b x^p f(x)dx$  by definition and the Taylor's expansion holds for complex arguments. You should have  $\phi_x(t) = e^{-t^2/2}$ .
- (c) Let  $y \sim N(\mu, \sigma^2)$ . Show  $\phi_y(t) = e^{it\mu - \sigma^2 t^2/2}$ . Hint: write  $y = \sigma x + \mu$ .

3. **Details of delta method proof.** Consider  $\mathbf{f} : \mathbb{R}^p \rightarrow \mathbb{R}^q$  where

$$\mathbf{f}(\mathbf{x}) = \begin{bmatrix} f_1(\mathbf{x}) \\ \vdots \\ f_q(\mathbf{x}) \end{bmatrix}.$$

For the  $i$ th component of  $\mathbf{f}$ , the multivariate Taylor's theorem gives us

$$f_i(\mathbf{x}) = f_i(\mathbf{x}_0) + [Df_i(\mathbf{x})|_{\mathbf{x}=\mathbf{x}_0}](\mathbf{x} - \mathbf{x}_0) + (\mathbf{x} - \mathbf{x}_0)' \mathbf{H}_i(\mathbf{x})(\mathbf{x} - \mathbf{x}_0),$$

where  $\lim_{\mathbf{x} \rightarrow \mathbf{x}_0} \mathbf{H}_i(\mathbf{x}) = \mathbf{0}$ . Here,  $Df_i(\mathbf{x})$  is the row-vector of first-partials  $[\frac{\partial f_i}{\partial x_1} \cdots \frac{\partial f_i}{\partial x_p}]$ .

- (a) Use Cauchy-Schwartz to show the absolute value of last term in the expansion is bounded by  $\|\mathbf{x} - \mathbf{x}_0\| \delta_i(\mathbf{x} - \mathbf{x}_0)$  where  $\lim_{\mathbf{x} \rightarrow \mathbf{x}_0} \delta_i(\mathbf{x} - \mathbf{x}_0) = 0$ .
  - (b) In the proof of the Delta method, further argue that  $\sqrt{n}\|\mathbf{t} - \boldsymbol{\mu}\| = O_p(1)$  and  $\|\boldsymbol{\delta}(\mathbf{t} - \boldsymbol{\mu})\| = o_p(1)$ .
4. Let  $y_1, \dots, y_n \stackrel{iid}{\sim} N(\boldsymbol{\mu}, \sigma^2)$ , i.e.  $\mathbf{y} \sim N_n(\boldsymbol{\mu}\mathbf{1}_n, \sigma^2\mathcal{I}_n)$ . Show  $\bar{y} = \frac{1}{n}\mathbf{1}'_n\mathbf{y}$  is indep. of  $s^2 = \frac{1}{n}\sum_{i=1}^n (y_i - \bar{y})^2$  using the property of multivariate normals where zero covariance implies independence. Do this by looking at  $\mathbf{A}\mathbf{y}$  where  $\mathbf{A} = \begin{bmatrix} \mathbf{1}'_n \\ \mathcal{I}_n - \frac{1}{n}\mathbf{1}_n\mathbf{1}'_n \end{bmatrix} \in \mathbb{R}^{(n+1) \times n}$  and making the appropriate argument.
5. Let  $\phi_{\mathbf{x}}(\mathbf{t})$  be the c.f. of  $\mathbf{x}$ . Find the c.f. of  $\mathbf{A}\mathbf{x} + \mathbf{b}$ , where  $\mathbf{A}$  and  $\mathbf{b}$  are conformable and fixed, in terms of  $\phi_{\mathbf{x}}(\mathbf{t})$ .
6. MKB 3.3.5.
7. MKB 4.2.3.