



<u>Neyman-Pearson Lemma</u>: If the likelihood ratio test that rejects H_0 in favor of H_A when

 $\lambda = \frac{f_0(\underline{x})}{f_A(\underline{x})} \leq c,$ has significance level α ,

then <u>any other</u> test having significance level at most α has power <u>less than or</u> <u>equal</u> to the power of the likelihood ratio test. (I.e., the LRT has <u>highest power</u> <u>among tests</u> with significance level α).

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Example: Consider a sample of size n from a normal distribution with variance 1.

Test $H_0:\mu=0$ vs. $H_A:\mu=1$ at $\alpha=0.05$.

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In some cases we can also show that the test is <u>uniformly most powerful</u> for a composite alternate hypotheses.

This happens if we can show it is most powerful for *every* simple alternate in H_A .

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Consider testing Test $H_0:\mu=0$ vs. $H_A:\mu>0$

and

Test $H_0:\mu=0$ vs. $H_A:\mu\neq 0$

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Example 2: Consider a binomial distribution with n=8 and unknown p. It is desired to test H_0 : p=0.2 versus H_A : p=0.4.

Confidence Intervals and Tests

There is a duality between confidence intervals and hypothesis tests. A confidence interval is found by "inverting" a two-sided test (and vice-versa).

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<u>Theorem A:</u> Suppose there is a test of level α for H₀: $\theta = \theta_0$, and let A(θ_0)=acceptance region

Then the set C={ θ : $\underline{X} \in A(\theta)$ } is a 100(1- α)% confidence region for θ .

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 $\frac{\text{Theorem B:}}{100(1-\alpha)\%} \text{ Let C}(\underline{X}) \text{ be a} \\ 100(1-\alpha)\% \text{ confidence region for } \\ \theta_0. \\$

Then A(θ_0)={X: $\theta_0 \in C(X)$ } is an acceptance region for a test of level α for H₀: θ = θ_0

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