

Chapter 16

Balanced Completely Random Design: $Y_{ij} = \beta_0 + \beta_1 X_{ij1} + \beta_2 X_{ij2} + \dots + \beta_{r-1} X_{ij,r-1} + \epsilon_{ij}$

- Instead of the predictors having their own distribution they are qualitative
- Their predicted coefficients are the mean responses.

Cell Means Model — denote r the number of levels of the factor with index i and $n_i = \#$ of cases for the i^{th} factor level w/ $n_T = \text{total } \#$ of cases ($n_T = \sum_i n_i$)

$Y_{ij} = \mu_i + \epsilon_{ij}$ w/ $E(Y_{ij}) = \mu_i$
 $\sigma^2(Y_{ij}) = \sigma^2(\mu_i + \epsilon_{ij}) = \sigma^2(\epsilon_{ij}) = \sigma^2$
 independent errors $\rightarrow Y_{ij} \overset{\text{ind}}{\sim} N(\mu_i, \sigma^2)$

Can be expressed in Matrix Terms

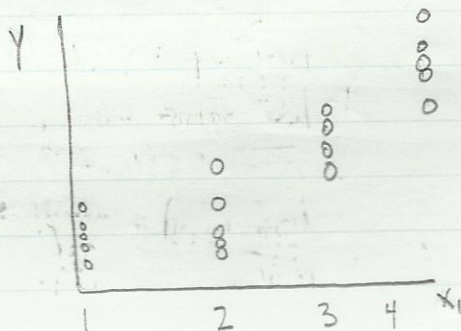
$Y = X\beta + \epsilon$

$$Y = \begin{bmatrix} Y_{11} \\ Y_{12} \\ Y_{21} \\ Y_{22} \\ Y_{31} \\ Y_{32} \end{bmatrix}, \quad X = \begin{bmatrix} 1 & 0 & 0 \\ 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \end{bmatrix}, \quad \beta = \begin{bmatrix} \mu_1 \\ \mu_2 \\ \mu_3 \end{bmatrix}, \quad \epsilon = \begin{bmatrix} \epsilon_{11} \\ \epsilon_{12} \\ \epsilon_{21} \\ \epsilon_{22} \\ \epsilon_{31} \\ \epsilon_{32} \end{bmatrix}$$

$E(Y) = [\mu_1, \mu_1, \mu_2, \mu_2, \mu_3, \mu_3]^T$ $\sigma^2(\epsilon) = \sigma^2 I$

Sas

```
proc sgscatter;
  plot Y * X1;
run;
```



SAS

```

proc glm;
  class x;
  model y = x;
  lsmeans x;

```

```

run;
proc ANOVA;
  class x;
  model y = x;
  means x;

```

run;

Levels of X	X LSMEAN
1	μ_1
2	μ_2
3	μ_3
⋮	⋮

Notation

$$Y_{ij} = \sum_{j=1}^{n_i} Y_{ij}$$

$$\bar{Y}_i = \frac{\sum_{j=1}^{n_i} Y_{ij}}{n_i} = \bar{Y}_i$$

$$\bar{Y}_{..} = \frac{\sum_{i=1}^r \sum_{j=1}^{n_i} Y_{ij}}{n_T} = \bar{Y}_{..}$$

$$\bar{Y}_{..} = \sum_{i=1}^r \sum_{j=1}^{n_i} Y_{ij}$$

$$\hat{\mu}_i = \bar{Y}_i$$

$$\hat{\mu}_{ij} = \bar{Y}_i$$

$$\sum_j e_{ij} = 0 \quad i=1, 2, \dots, r \quad \text{errors for each factor level sum to 0}$$

Given by both proc GLM and ANOVA above

$$SSTO = \sum_i \sum_j (Y_{ij} - \bar{Y}_{..})^2$$

$$SSTreatment = \sum_i n_i (\bar{Y}_i - \bar{Y}_{..})^2$$

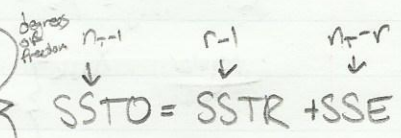
$$SSE = \sum_i \sum_j (Y_{ij} - \bar{Y}_i)^2 = \sum_i \sum_j e_{ij}^2$$

$$MStreatment = \frac{SSTreatment}{r-1}$$

$$MSE = \frac{SSE}{n_T - r}$$

$$E(MStreatment) = \sigma^2 + \frac{\sum_i n_i (\mu_i - \mu_0)^2}{r-1}$$

$$E(MSE) = \sigma^2$$



F Test for equality of factor level means

$$H_0: \mu_1 = \mu_2 = \dots = \mu_r$$

H_a : not all μ_i are equal

$$F^* = \frac{MSTreatment}{MSE}$$

Reject when $F^* > F(1-\alpha; r-1, n-r)$

*p-value for this given by proc anova and GLM on p 69

Factor Effects Model $Y_{ij} = \mu_0 + \tau_i + \epsilon_{ij}$ $\tau_i = \mu_i - \mu$

\uparrow \uparrow
mean factor
response

$i = 1, \dots, r$
 $j = 1, \dots, n_i$
 $\sum_i \tau_i = 0$

Weighted Mean $\mu_0 = \sum_{i=1}^r w_i \mu_i$ $\sum_{i=1}^r w_i = 1$
 $\sum_{i=1}^r w_i \tau_i = 0$

Test for equality of factor level means

$$H_0: \tau_1 = \tau_2 = \dots = \tau_r = 0$$

H_a : not all = 0

use same F test above.

Factor Effects Model

$$\sum_i \tau_i = 0 \rightarrow \tau_r = -\tau_1 - \tau_2 - \dots - \tau_{r-1}$$

We can fit the linear model w/ $\mu, \tau_1, \tau_2, \dots, \tau_{r-1}$

$$\text{So: } Y_{ij} = \mu_0 + \tau_1 X_{ij1} + \dots + \tau_{r-1} X_{ij,r-1} + \epsilon_{ij}$$

regular $X_{ijk} = \begin{cases} 1 & \text{if case is from factor level } k \\ -1 & \text{if case is from factor level } r \\ 0 & \text{ow} \end{cases}$

weighted $X_{ijk} = \begin{cases} 1 & \text{if case is from factor level } k \\ -\frac{n_k}{n_r} & \text{if case is from factor level } r \\ 0 & \text{ow} \end{cases}$

* Follow same linear regression from 704

SAS

/* set up taus in data step */

proc glm;

model y = tau1 tau2 ...;

estimate 'tau r' tau1 -1 tau2 -1 ... tau r -1 -1;

run;